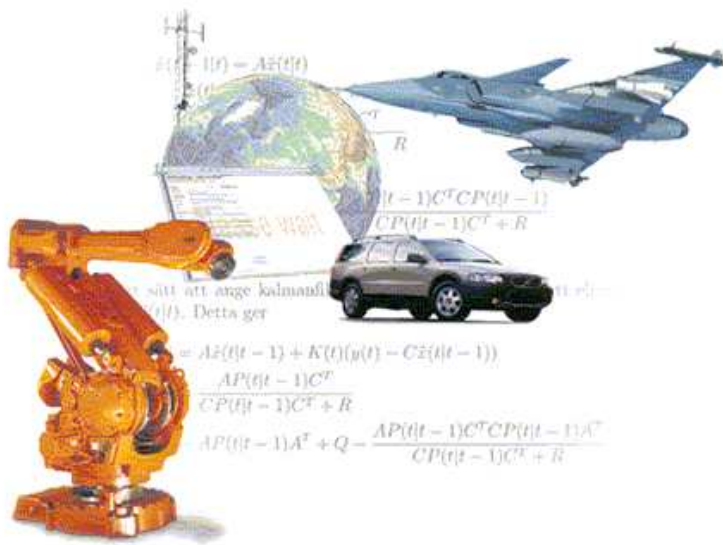


# A (Simplistic) Perspective on Nonlinear System Identification

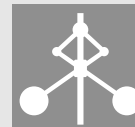


Lennart Ljung

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Abstract: Nonlinear System Identification is really curve fitting

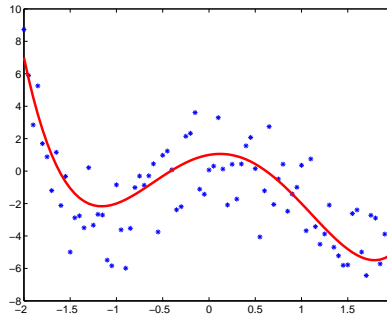


## Abstract: Nonlinear System Identification is really curve fitting

1. The basic questions and (statistical) tools illustrated for a simple curve fitting problem.
2. Nonlinear dynamical models: Parameterizations, problems and techniques.



Most basic ideas from system identification, choice of model structures and model sizes are brought out by considering the basic curve fitting problem from elementary statistics.



Unknown function  $g_0(x)$ . For a sequence of  $x$ -values (regressors)  $\{x_1, x_2, \dots, x_N\}$  (that may or may not be chosen by the user) observe the corresponding function values with some noise:

$$y(k) = g_0(x_k) + e(k)$$

Construct an estimate  $\hat{g}_N(x)$  from  $\{y(1), x_1, y(2), x_2, \dots, y(N), x_N\}$

.



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The error  $\hat{g}_N(x) - g_0(x)$  should be “as small as possible”

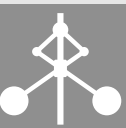
Approaches:

- **Parametric:** Construct  $\hat{g}_N(x)$  by searching over a parameterized set of candidate functions.
- **Non-parametric:** Construct  $\hat{g}_N(x)$  by smoothing over (carefully chosen subsets of)  $y(k)$



Search for the function  $g_0$  in a parameterized family of functions:

$$g(x, \theta) = \sum_{k=1}^n \alpha_k f_k(x, \tilde{\theta}_k), \quad \theta = \{\alpha_k, \tilde{\theta}_k, k = 1, \dots, n\}$$



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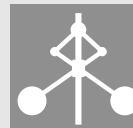
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Examples:

**Polynomial:**  $g(x, \theta) = \theta_1 + \theta_2 x + \dots + \theta_n x^{n-1}$

**Piecewise constant:**  $g(x, \theta) = \sum_{k=1}^n \alpha_k U(\beta_k(x - \gamma_k)),$

$U(x)$  is the unit pulse.



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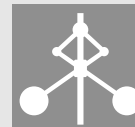
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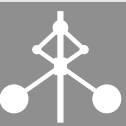
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The basic form is

$$g(x, \theta) = \sum_{k=1}^N \alpha_k \kappa(\beta_k(x - \gamma_k))$$



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- Wavenets
- Neuro-Fuzzy



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$$y(t) = g_0(x_t) + e(t)$$

Least Squares:

$$\hat{\theta}_N = \arg \min_{\theta} V_N(\theta)$$

$$V_N(\theta) = \frac{1}{N} \sum_{t=1}^N |y(t) - g(x_t, \theta)|^2$$



$$y(t) = g_0(x_t) + e(t)$$

Weighted Least Squares:

$$\hat{\theta}_N = \arg \min_{\theta} V_N(\theta)$$

$$V_N(\theta) = \frac{1}{N} \sum_{t=1}^N |y(t) - g(x_t, \theta)|^2 / \lambda_t$$

$\lambda_t$  Proportional to 'reliability' of  $t$ :th measurement  $\sim Ee^2(t)$



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Weighted Least Squares:

$$\hat{\theta}_N = \arg \min_{\theta} V_N(\theta)$$

$$V_N(\theta) = \frac{1}{N} \sum_{t=1}^N L(x_t) |y(t) - g(x_t, \theta)|^2 / \lambda_t$$

$\lambda_t$  Proportional to 'reliability' of  $t$ :th measurement  $\sim Ee^2(t)$

A extra weighting  $L(x_t)$  could also reflect the 'relevance' of the point  $x_t$ .

('Focus in fit')





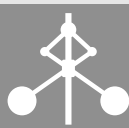
$$y(t) = g_0(x_t) + e(t)$$

(Regularized) Least squares:

$$\hat{\theta}_N = \arg \min_{\theta} V_N(\theta) + \delta |\theta|^2$$

$$V_N(\theta) = \frac{1}{N} \sum_{t=1}^N |y(t) - g(x_t, \theta)|^2$$

$\delta |\theta|^2$  penalizes excessive model flexibility. Could come in various forms.



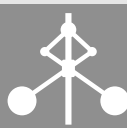
$$\hat{\theta}_N = \arg \min_{\theta} \frac{1}{N} \sum_{t=1}^N \ell(y(t) - g(x_t, \theta), t)$$

- Maximum likelihood  $\ell(z) = -\log p(z)$
- “unknown-but-bounded”:  $\min_{\theta} \max_t |y(t) - g(x_t, \theta)|$
- ‘Support vector machines’:  $\min \sum |y(t) - g(x_t, \theta)|_{\epsilon}$  ( $\epsilon$ -insensitive  $L_1$  norm)

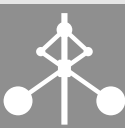
Regularization by

$$V_N(\theta) + \delta|\theta| \quad \text{or} \quad \min V_N(\theta), |\theta| < C$$

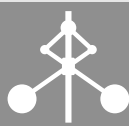
LARS, LASSO, nn-garotte ...



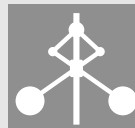
So, the choice of parameters within a parameterized model is not that difficult: Fit to the observed data, by one criterion or another.  
The choice of model size and model parameterization is a more interesting issue.



Except for very simple parameterizations  $g(x, \theta)$ , the distribution of  $\hat{\theta}_N$  cannot be calculated (mainly due to “arg min”).

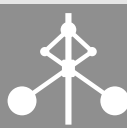


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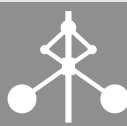
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- $H(\theta) = \lim_{N \rightarrow \infty} H_N(\theta) = EL(x_t) |g_0(x_t) - g(x_t, \theta)|^2 / \lambda_t$
- Main Result:  $\lim_{N \rightarrow \infty} \hat{\theta}_N = \theta^* = \arg \min H(\theta)$

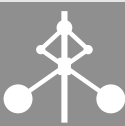


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- Main Result:  $\lim_{N \rightarrow \infty} \hat{\theta}_N = \theta^* = \arg \min H(\theta)$
- The asymptotic distribution of  $\sqrt{N}(\hat{\theta}_N - \theta^*)$  is normal with zero mean and covariance matrix  $P = \lambda [E\psi(t)\psi^T(t)]^{-1}$ ,  $\psi(t) = \frac{d}{d\theta} g(x_t, \theta^*)$
- **“Cov  $\hat{\theta}_N \sim \frac{\lambda}{N} [E\psi(t)\psi^T(t)]^{-1}$ ”** (Decreases with more regularization)



- Effective number of parameters (depending on parameter dimension and regularization) is a trade-off between bias and variance
- This trade-off is favored by grey-box models and by adaptive choices of basis functions for the parameterization

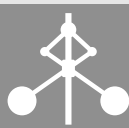




A simple idea is to locally smooth the noisy observations of the function values:

$$\hat{g}_N(x) = \sum_{k=1}^N C(x, x_k) y(k)$$

$$\sum_{k=1}^N C(x, x_k) = 1 \quad \forall x$$

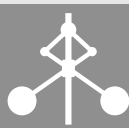


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Often  $C(x, x_k) = \tilde{c}(x - x_k) / \lambda_k$  and  $\tilde{c}(r) = 0$  for  $|r| > \beta$ ,  $\beta =$  the “bandwidth”  
These are known as “kernel methods” in statistics.



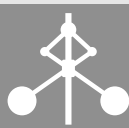
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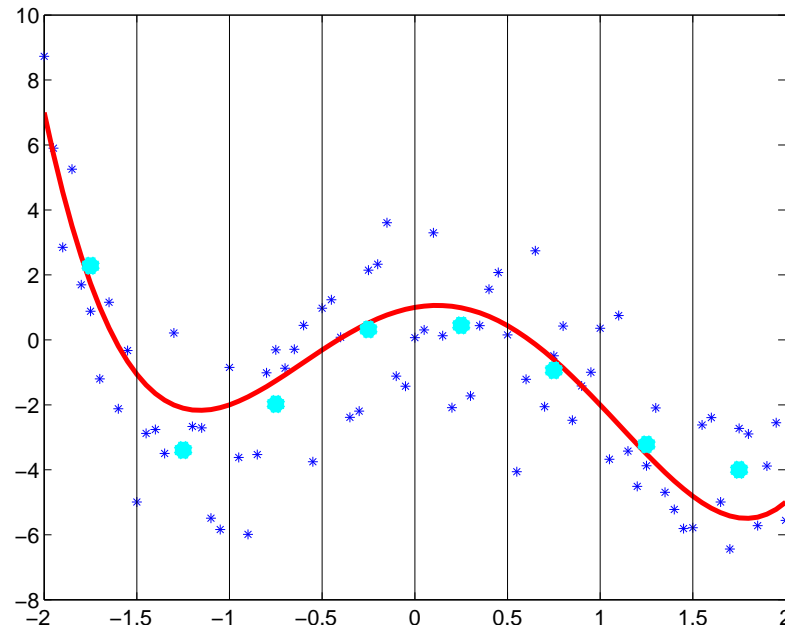
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If  $C(x, x_t)$  is chosen so that it is non-zero ( $= 1/k$ ) only for  $k$  observed values  $x_t$  around  $x$ , this is the **k-nearest neighbor method**.

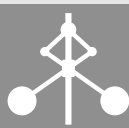




$C(x, x_k) = U((x - x_k)/\beta)$ ;  $U(\cdot)$  the unit pulse.  $\beta = 0.25$ .

Cyan dots: Computed for  $x = -1.75 : 0.5 : 1.75$

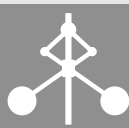
Bias-Variance Trade-off: ...



- Local polynomial models
  - Adjust polynomials in local neighborhoods around  $x$ , Evaluate them in  $x$ .
- Direct weight optimization

$$\hat{g}_N(x) = \sum w_k(x)y(k), \quad \text{Choose } \{w_k\} \text{ for each } x$$

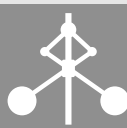
- Typically “Model-on-Demand” rather than “Off-the-Shelf”



Data: outputs and inputs

$$\{y(1), u(1), \dots, y(N), u(N)\} = Z^N$$

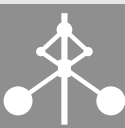
- General aspects
- Black-box models
- Light-Grey-box models
- Dark-Grey-box models



A mathematical model for the system is a function from the past input-output data to the space where the output at time  $t$ ,  $y(t)$  lives, in general

$$\hat{y}(t|t-1) = g(Z^{t-1}, t)$$

The function can be thought of as a predictor of the next output.



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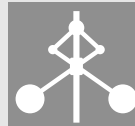
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Let us split it into one mapping from  $Z^{t-1}$  to a regression vector  $\varphi(t)$  of fixed dimension  $d$  and a mapping  $g$  from  $R^d$  to  $R$ :

$$g(Z^{t-1}, t) = g(\varphi(t))$$

$$\varphi(t) = \varphi(Z^{t-1}, t) \quad \text{Finding } \varphi(t) \text{ could itself be an estimation problem}$$





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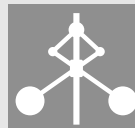
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Leaves two problems:

1. Choose the mapping  $g(\varphi)$  – **Same as in curvefitting**
2. Choose the regression vector  $\varphi(t)$  – **“State”**

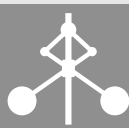


Suppose  $\varphi(t) = [y(t-1), u(t-1)]^T$

The (one-step ahead) **predicted** output at time  $t$  for a given model  $\theta$  is then

$$\hat{y}_p(t|\theta) = g([y(t-1), u(t-1)]^T, \theta)$$

It uses the previous measurement  $y(t-1)$ .



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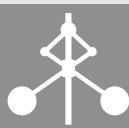
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A tougher test is to check how the model would behave in simulation, i.e. when only the input sequence  $u$  is used. The **simulated** output is obtained as above, by replacing the measured output by the simulated output from the previous step:

$$\hat{y}_s(t, \theta) = g([\hat{y}_s(t-1, \theta), u(t-1)]^T, \theta)$$



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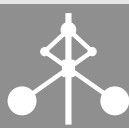
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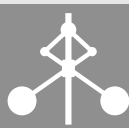
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**Notice a possible stability problem!**



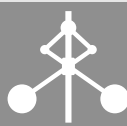
- Black
  - Parametric – Non-Parametric: see Curve Fitting
- Light-Grey
  - Physical modeling
- Dark-Grey
  - Semi-physical modeling
  - Block-oriented models
  - Local linear models and their cousins



Perform physical modeling (e.g. in MODELICA) and denote unknown physical parameters by  $\theta$ . Collect the model equations as

$$\dot{x}(t) = f(x(t), u(t), \theta)$$

$$y(t) = h(x(t), u(t), \theta)$$

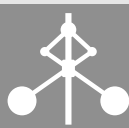


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(or in DAE, Differential Algebraic Equations, form.)



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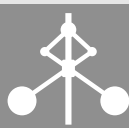
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(or in DAE, Differential Algebraic Equations, form.) For each parameter  $\theta$  this defines a simulated (predicted) output  $\hat{y}(t|\theta)$  which is the parameterized function

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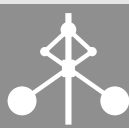
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in somewhat implicit form. To be a correct predictor this really assumes white measurement noise. Some more sophisticated noise modeling is possible, usually involving *ad hoc* non-linear observers.



Perform physical modeling (e.g. in MODELICA) and denote unknown physical parameters by  $\theta$ . Collect the model equations as

$$\dot{x}(t) = f(x(t), u(t), \theta)$$

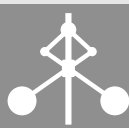
$$y(t) = h(x(t), u(t), \theta)$$

(or in DAE, Differential Algebraic Equations, form.) For each parameter  $\theta$  this defines a simulated (predicted) output  $\hat{y}(t|\theta)$  which is the parameterized function

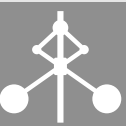
$$\hat{y}(t|\theta) = g(Z^{t-1}, \theta)$$

in somewhat implicit form. To be a correct predictor this really assumes white measurement noise. Some more sophisticated noise modeling is possible, usually involving *ad hoc* non-linear observers.

The approach is conceptually simple, but could be very demanding in practice.

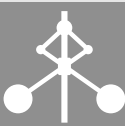


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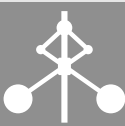
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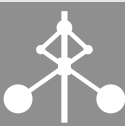
Simple examples: . . . .



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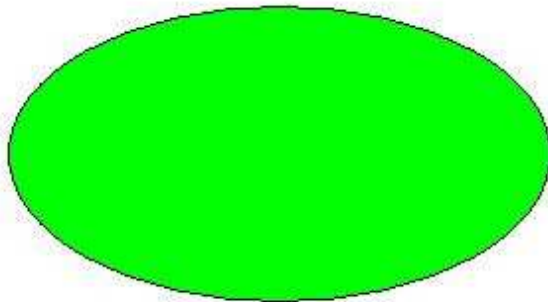
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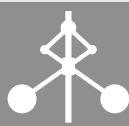
Building Blocks:

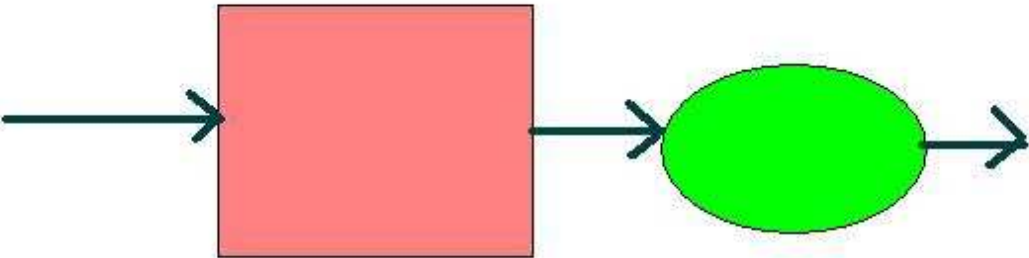


Linear Dynamic System  
 $G(s)$

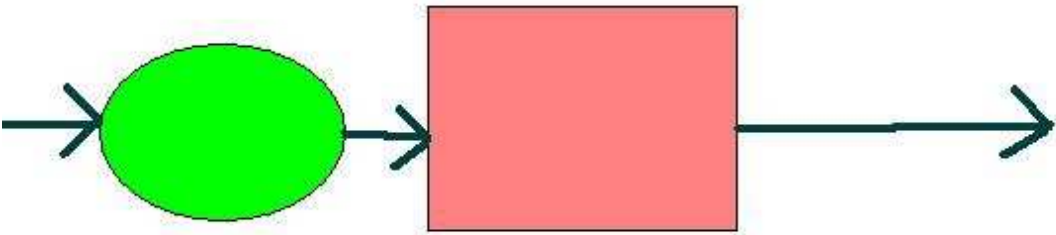


Nonlinear static function  
 $f(u)$

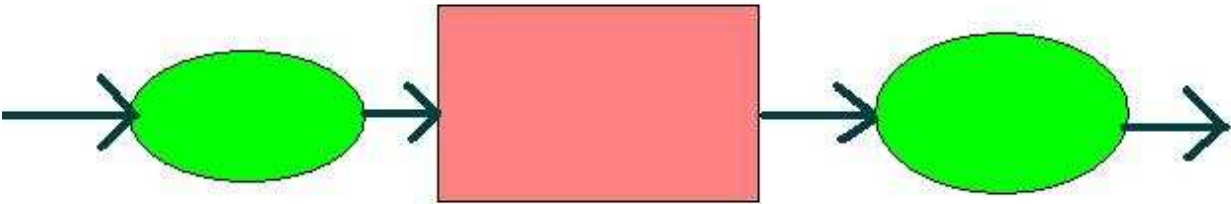




Wiener



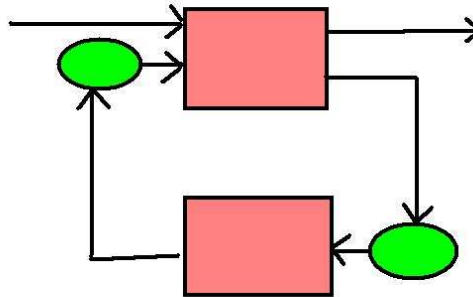
Hammerstein



Hammerstein-  
Wiener





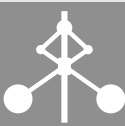


With the linear blocks parameterized as a linear dynamic system and the static blocks parameterized as a function (“curve”), this gives a parameterization of the output as

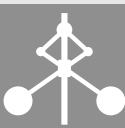
$$\hat{y}(t|\theta) = g(Z^{t-1}, \theta)$$

and the general approach of model fitting can be applied.

However, in this contexts many algorithmic variants have been suggested.



Non-linear systems are often handled by linearization around a working point. The idea behind **Local Linear Models** is to deal with the nonlinearities by selecting or averaging over relevant linearized models.

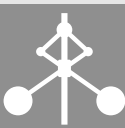


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Let the measured working point variable be denoted by  $\rho(t)$  (sometimes called **regime variable**). If the regime variable is partitioned into  $d$  values  $\rho_k$ , the predicted output will be

$$\hat{y}(t) = \sum_{k=1}^d w_k(\rho(t), \rho_k, \eta) \hat{y}^{(k)}(t)$$

where  $\eta$  is a parameter that describes the partitioning

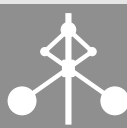


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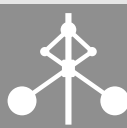
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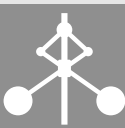
If the prediction  $\hat{y}^{(k)}(t)$  corresponding to  $\rho_k$  is linear in the parameters,  $\hat{y}^{(k)}(t) = \varphi^T(t)\theta^{(k)}$  the whole model will be a linear regression for a fixed  $\eta$ .



The model structure

$$\hat{y}(t, \theta, \eta) = \sum_{k=1}^d w_k(\rho(t), \eta) \varphi^T(t) \theta^{(k)}$$

is also an example of a **hybrid** model (piecewise linear). If the partition is to be estimated too, the problem is considerably more difficult.



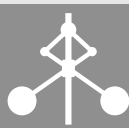
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**Linear Parameter Varying (LPV)** models are also closely related:

$$\begin{aligned} \dot{x} &= A(\rho(t))x + B(\rho(t))u \\ y &= C(\rho(t))x + D(\rho(t))u \end{aligned}$$



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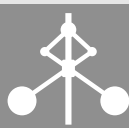
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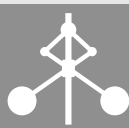
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**Notice the link to non-parametric Local Polynomial Models in statistics!**

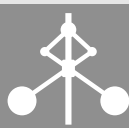




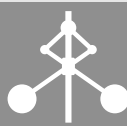
- A nonlinear model can be seen as nonlinear mapping from past data to the space where the output lives:  $\hat{y}(t|t-1) = g(Z^{t-1}, t)$ . Observations are then  $y(t) = \hat{y}(t|t-1) + e(t)$ .



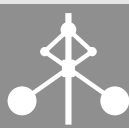
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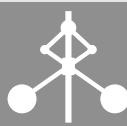
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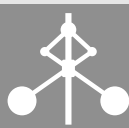
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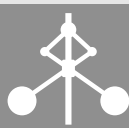
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- Non-convexity of the optimization remains one of the more serious problems for most parametric methods.







- For Tomorrow's Panel Discussion ...

